

Omitted Variable Bias



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Outline

- Endogeneity, Confounders, and Omitted Variable Bias
- Instrumental Variable (IV) Analysis and OVB
- Impact threshold of a confounding variable (ITCV)
- Concluding Thoughts and Recommendations



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Capturing Causality in Organizational Research



Causality represents a challenge for organizational researchers



For example, the strategy field is focused on complex, interrelated factors that influence competitive advantage and firm performance (Bowman, Singh, & Thomas, 2002; Nag, Hambrick, & Chen, 2007)



When examining these relationships, scholars often rely almost exclusively on non-experimental design (Bergh et al., 2004; Bettis et al., 2014; Hamilton & Nickerson, 2003)



Because of this reality, researchers faces several empirical and analytical challenges



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Casual Inference and Omitted Variable Bias

“Yes, but have you controlled for...”

(Frank, 2000: 149)




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Endogeneity

- Endogeneity occurs when an independent variable is correlated with the error term in a statistical model
- The basic OLS regression model is:

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$


$$Cov(x_i, \varepsilon_i) \neq 0$$

- Endogeneity may derive from:
 - Omitted variable
 - Measurement error
 - Simultaneous causality
 - Sample Selection



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Campbell's Threats to Validity in Econometric Terms

Threats to internal validity	Threats to external validity
<ul style="list-style-type: none">• Omitted variables• Trends in outcomes• Misspecified variances• Mismeasurement• Political economy• Simultaneity• Selection• Attrition• Omitted interactions	<ul style="list-style-type: none">• Interaction of selection and treatment• Interaction of setting and treatment• Interaction of history and treatment <p>(Meyer, 1995)</p>



Omitted Variable Bias and Confounding

- Omitted variable bias (OVB) arises from not including a relevant variable that belongs in the population model
- OV B = excluding a relevant variable = underspecifying the model = short regression = left out variables error (LOVE)
- Confounding is the bias caused by common causes of a treatment and outcome
 - Produces “spurious correlation” and biased results
- In observational studies, the goal is to avoid confounding
- Pervasive in the organizational research:
 - Director expertise on monitoring (confounding: motivation)
 - CEO influence on firm performance (confounding: managerial ability)
 - Resources and firm performance (confounding: external environment)
- No unmeasured confounding assumes that we’ve measured all sources of confounding.



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Omitting ability when examining the effect of education on wages

- If the true model in the population is:

$$wages = \beta_0 + \beta_1 education + \beta_2 ability + \mu$$

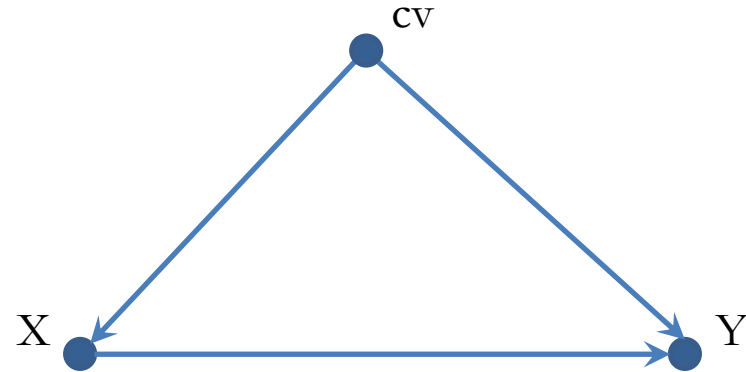
- Using a sample, what happens if we only model:

$$wages = \beta_0 + \beta_1 education_i + \mu_i$$

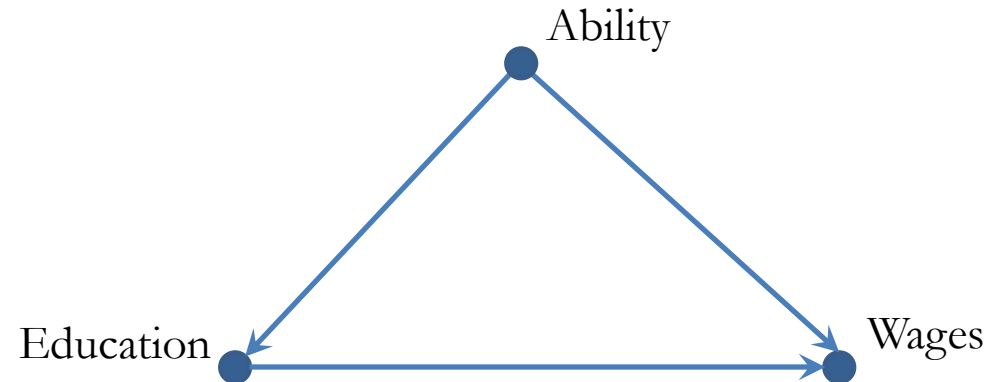
- OVB describes the difference in regression estimates between these two equations (Angrist & Pischke, 2008)



Confounding Variables in DAGs



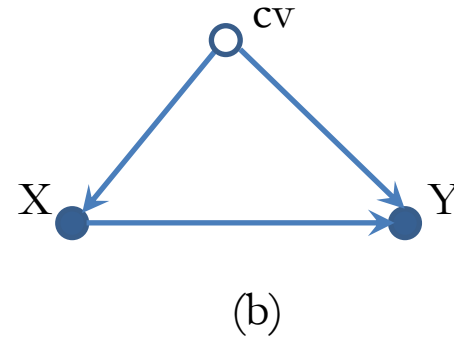
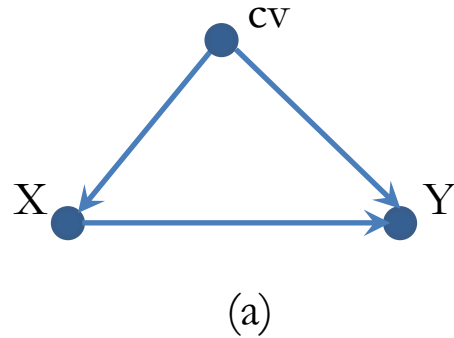
C is a confounder of the proposed causal relationship between X and Y



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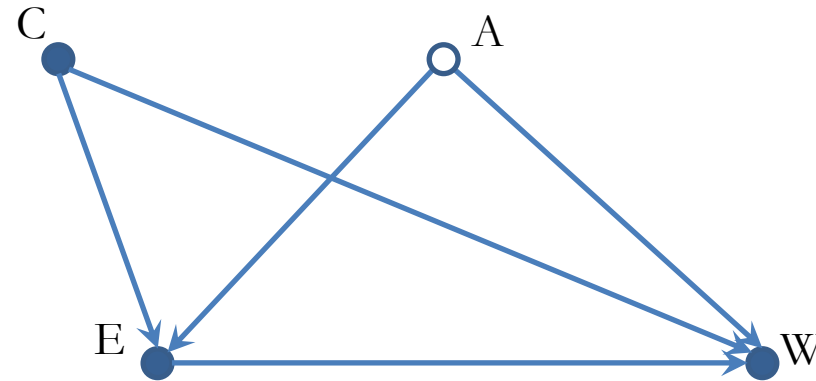
Two Graphs in which the Causal Effect of X on Y is Confounded by C



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Expanding the DAG for Wages



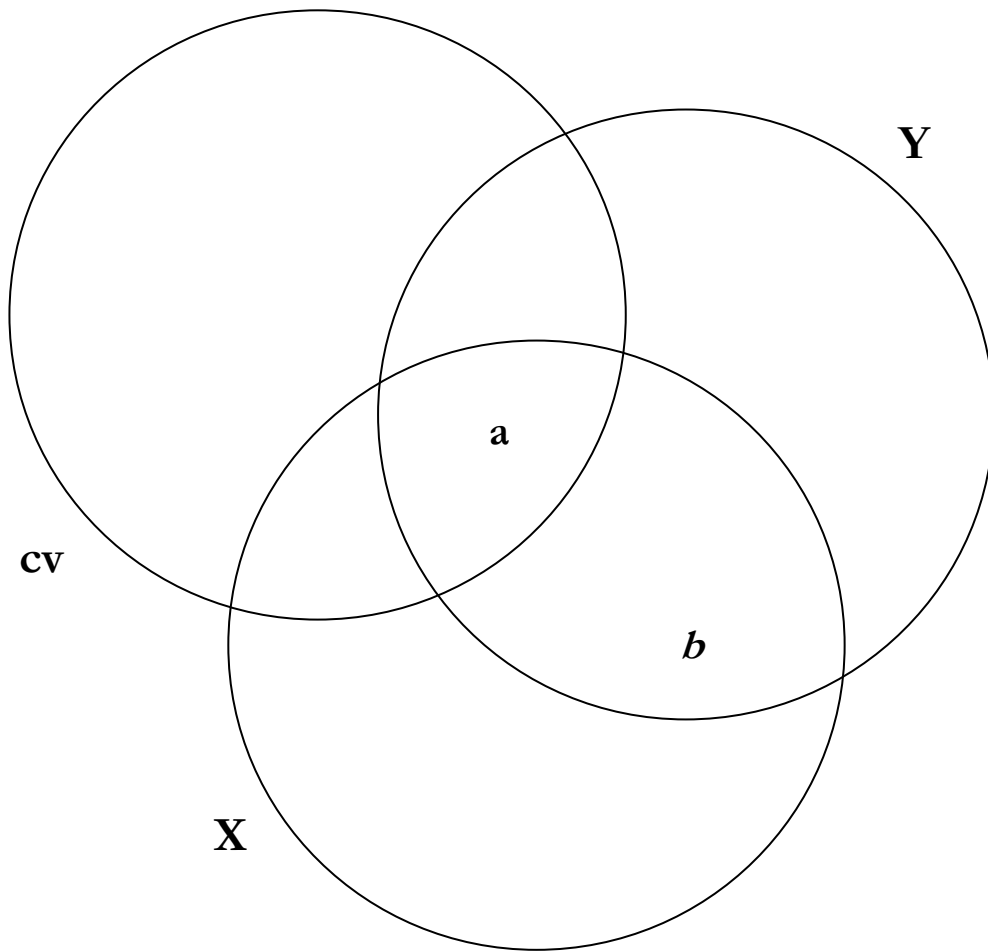
$$Wages = \beta_0 + \beta_1 education + \beta_2 C + \beta_3 ability + \mu$$



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A Ballantine Illustration of Education on Wages



- Back to the short regression equation:

$$wages = \beta_0 + \beta_1 education + \mu$$

- Circles Y and X represent variation *wages* and *education*, respectively
- The area **a** + **b** represent the overlap in variation between Y and X
- However, the area **a** represent overlap in variation between X (*education*) and the omitted variable *ability* and Y(*wage*), which creates a correlation between the error term (μ) and X
- If Y were regressed on X, the info in the a + b would be used to estimate $\beta_1 education$.



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Example: Omitting Ability when Examining the Effect of Education on Wages

$$wages = \beta_0 + \beta_1 education + \beta_2 ability + \mu$$

$$ability = \delta_0 + \delta_1 education + \vartheta$$

Will both be positive

$$wages = (\beta_0 + \beta_2 \delta_0) + (\beta_1 + \beta_2 \delta_1) + (\beta_2 \vartheta + \mu)$$

The return to wages β_1 will be overestimated because $\beta_2 \delta_1 > 0$. It will look as if individuals with more education earn very high wages, but this is partly due to the fact that these individuals with more education also have greater ability on average.



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Summary of Bias in $\widetilde{\beta}_1$

If the model in the population is:

$$y_i = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \varepsilon_i$$

	$Cov(x_1, x_2) > 0$	$Cov(x_1, x_2) < 0$
$\beta_2 > 0$	Positive Bias	Negative Bias
$\beta_2 < 0$	Negative Bias	Positive Bias



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When is OVB Not an Issue?

If the model in the population is:

$$wages = \beta_0 + \beta_1 education + \beta_2 ability + \varepsilon_i$$

1. If $\beta_2 = 0$ in the population model, $\widetilde{\beta}_1$ is unbiased
2. If education and ability were uncorrelated, $\widetilde{\beta}_1$ is unbiased

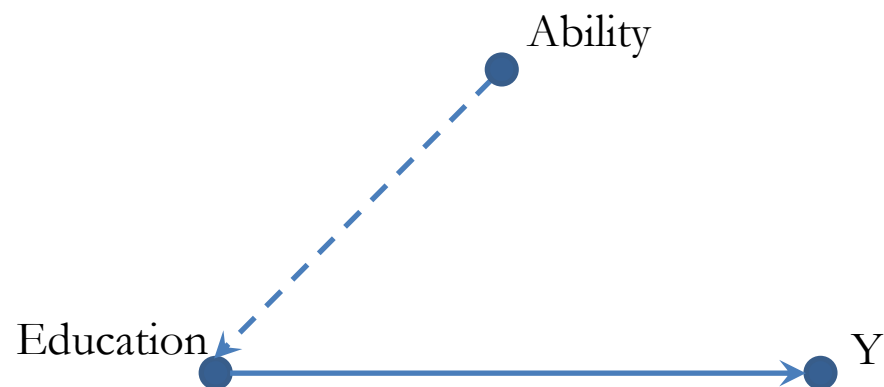


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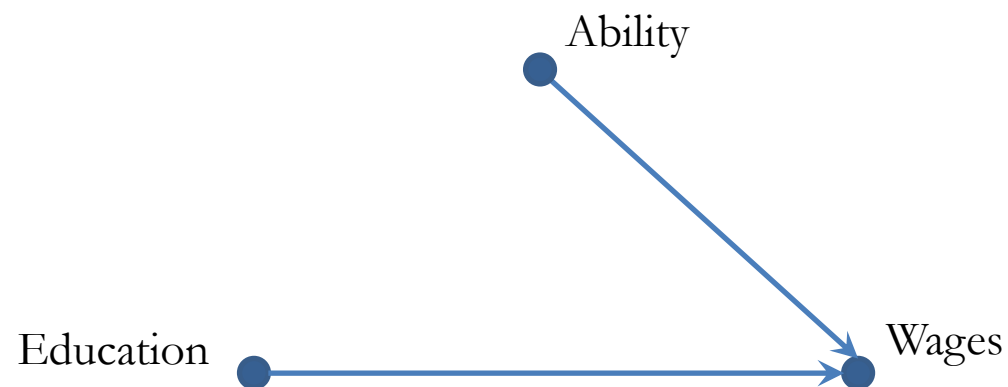
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When is OVB Not an Issue?

Condition 1: $\beta_2 = 0$ in the population model



Condition 2: Education and Ability are uncorrelated



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Instrumental Variable (IV) Analysis and OVB



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OVV and IV Analysis



Concern about OVB is often a key motivating reason for adopting instrumental variable techniques



These techniques typically involve a two-step procedure



While these techniques can help alleviate the OVB concern, they also have critical assumptions that must be met



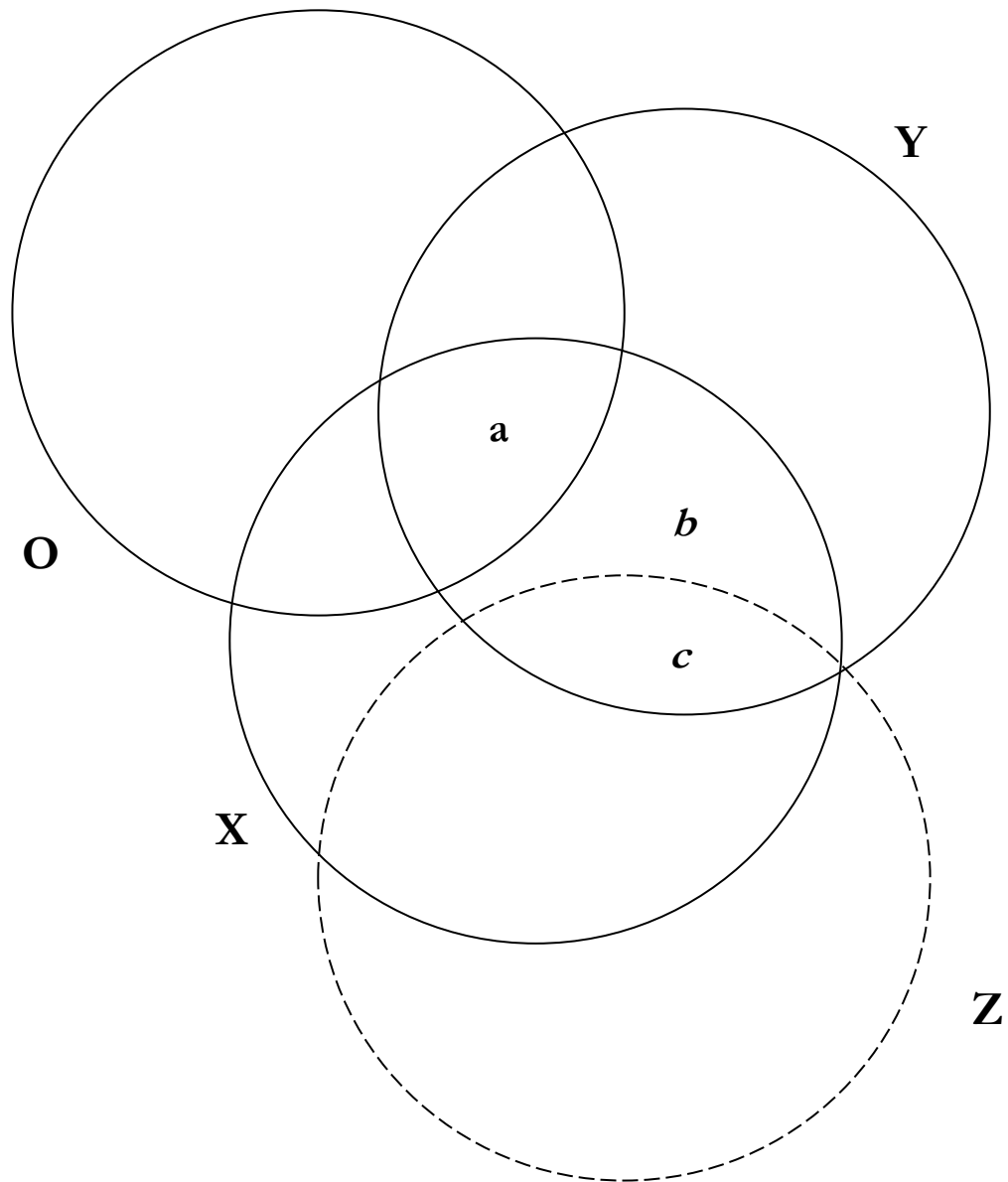
Even when these assumptions are met, instrumental variable techniques are often less efficient than OLS regression



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A Ballantine Illustration of IV Analysis Logic



- The Z variable represents an IV
- Suppose X is regressed on Z. The predicted area, \hat{X} , is represented by the c area
- Now regress Y on \hat{X} to produce an estimate of β_1 *education*
- In this case, area c is only used to form the estimate
- Since area c corresponds to variation in Y arising from variation in X, the resulting estimate of β_1 *education* is unbiased

An IV Approach to Education on Wages

$$y = \alpha + \beta_1 x_1 + \beta_2 x_2 \cdots + \beta_{k-1} x_{k-1} + \varepsilon_i$$

Education variable that is suspected to be endogenous

First stage (= reduced form regression):

The endogenous explanatory variable x_1 is predicted using only exogenous information

$$\hat{x}_1 = \hat{\alpha} + \hat{\alpha}_1 x_2 \cdots + \hat{\alpha}_{k-1} x_{k-1} + \hat{\alpha}_k z_k$$

Exogenous variable = distance to a 4-year college

Second stage (= OLS with x_1 replaced by its prediction from the first stage):

$$y = \alpha + \beta_1 \hat{x}_1 + \beta_2 x_2 \cdots + \beta_{k-1} x_{k-1} + \varepsilon_i$$

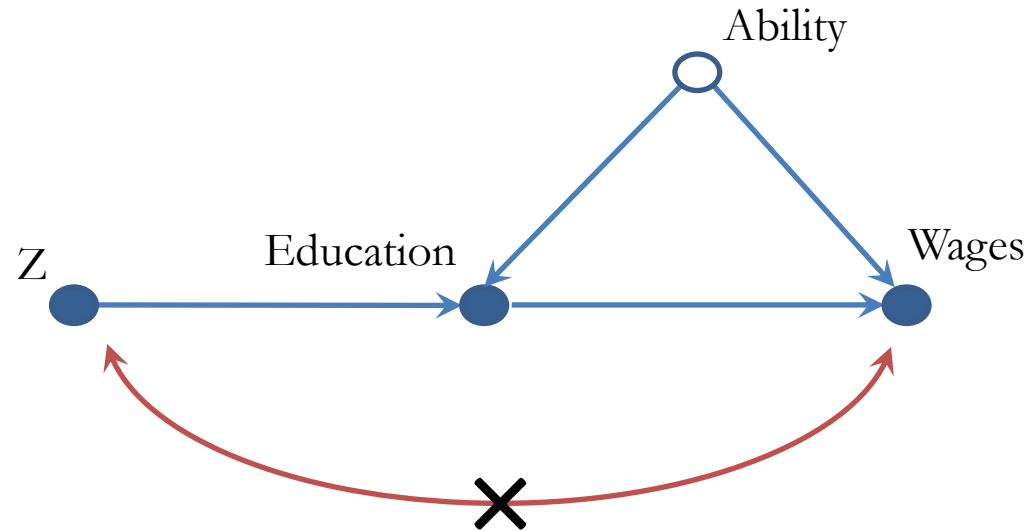
Intuition: The **predicted value** is the value of the endogenous variable as a function of the exogenous instrument. This isolates the variance in the endogenous variable that is exogenous.



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Basic IV Setup with DAGs



- Z is the instrument, *education* is the treatment, and *ability* is the unmeasured confounder
- Exclusion restriction
 - No common causes of the instrument and the outcome
 - No direct or indirect effect of the instrument on the outcome not through the *education*.
- First-stage relationship: $Z \rightarrow \text{education}$



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Finding Instrumental Variables

Requirements:

- Excluded from the regression model
- Relevance (strong correlation with potential endogenous variable)
 - Multiple strong instruments optimal
- Exogeneity (NOT correlated with error)



Relevance: How Do We Know?

- In the first stage regression (where endogenous IV is the DV), the instrument(s) should explain a significant portion of the DV.
- F-statistics can be used to examine this assumption.



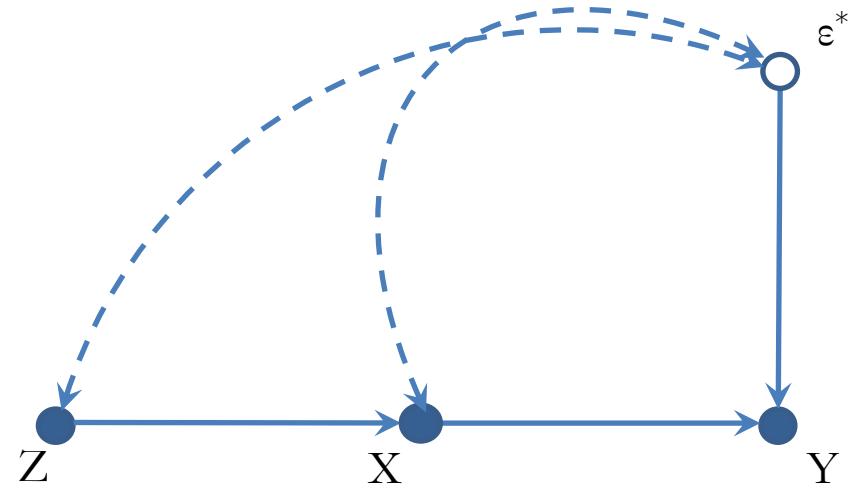
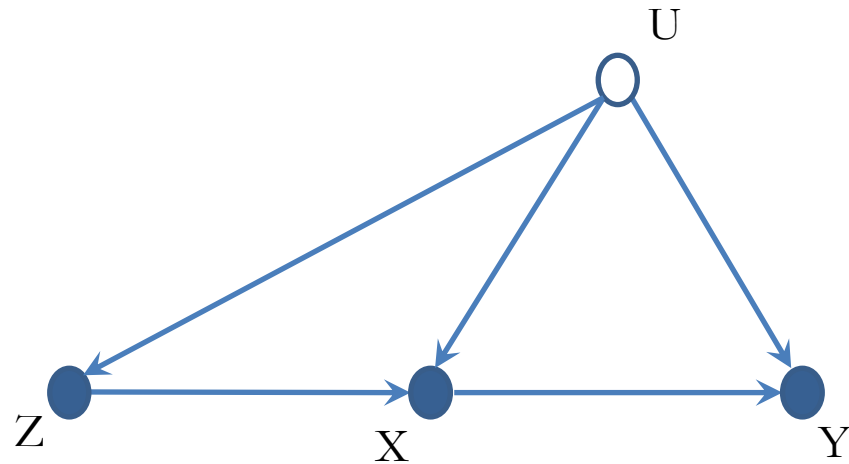
Instrumental Variables: Relevance

Table 1. Main findings OLS vs. single instrument

	OLS	Weak inst.	Moderate inst.	Weak and endo instr.	Moderate and endo instr.
PANEL A: Sample size: 500, True B : 0.1					
No endogeneity					
Beta	0.100	0.088	0.100	1.095	0.399
SE	0.045	0.469	0.138	0.633	0.141
95% interval	95%	100%	96%	84%	43%
% significant	60%	1%	9%	26%	84%
Low endogeneity (0.1)					
Beta	0.200	0.076	0.107	1.073	0.406
SE	0.045	0.478	0.135	0.598	0.139
95% interval	39%	100%	96%	78%	40%
% significant	69%	1%	12%	34%	86%
Medium endogeneity (0.3)					
Beta	0.430	0.097	0.092	1.142	0.404
SE	0.042	0.475	0.136	0.556	0.129
95% interval	0%	100%	95%	66%	36%
% significant	100%	4%	13%	45%	86%
PANEL B: Sample size: 500, True B : 0					
No endogeneity					
Beta	0.001	-0.009	-0.004	0.940	0.335
SE	0.045	0.474	0.137	0.610	0.157
95% interval	95%	100%	95%	84%	44%
% significant	5%	1%	5%	16%	56%
Low endogeneity (0.1)					
Beta	0.100	0.051	-0.009	1.016	0.336
SE	0.044	0.489	0.136	0.608	0.154
95% interval	39%	100%	95%	77%	40%
% significant	61%	0%	5%	23%	60%
Medium endogeneity (0.3)					
Beta	0.330	-0.001	-0.007	0.987	0.338
SE	0.042	0.466	0.135	0.536	0.143
95% interval	0%	98%	95%	63%	35%
% significant	100%	2%	5%	37%	65%



Exogeneity Assumption Violated in DAGs



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Tests for Exogeneity

- Testing for exogeneity requires that the number of instruments exceeds the number of endogenous regressors (i.e., the equation is “overidentified”).

Basile (2008) summarizes three tests (and provides Stata code):

- The Sargan or Hansen J-statistic
- The Bassmann statistic
- The difference-in-Sargan statistic

*Note: a *failure* to reject each type of statistic means that the instruments can be considered exogenous.



Instrumental Variables: Exogeneity

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Where Do You Find Good IVs?

- Finding IVs that meet the relevance and exogeneity requirements is challenging
- This is the major obstacle with IV techniques
- If you only have one IV, you can never fully know if the criteria are met

“All instruments arrive on the scene with a dark cloud of invalidity hanging overhead. This cloud never goes entirely away, but researchers should chase away as much of the cloud as they can.”

Murray (2006: 114)



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Impact threshold of a confounding variable (ITCV)



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Quantifying the OVB Problem

Impact threshold of a confounding variable (ITCV)

What is the minimum correlation between a confounder variable and the independent variable/dependent variable for OVB to have created a significant effect where none really exists?



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Causal Inferences

- Frank and colleagues suggest that you can make a causal inference from an observational study
- However, you just might be wrong
- The focus shifts from bias to changes in the causal inference
- But what would it take for the inference to be wrong?



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Key steps

1

Establish
Correlation
Between
predictor of
interest and
outcome

2

Define a
Threshold for
Inference

3

Calculate the
Impact Necessary
to Invalidate the
Inference

4

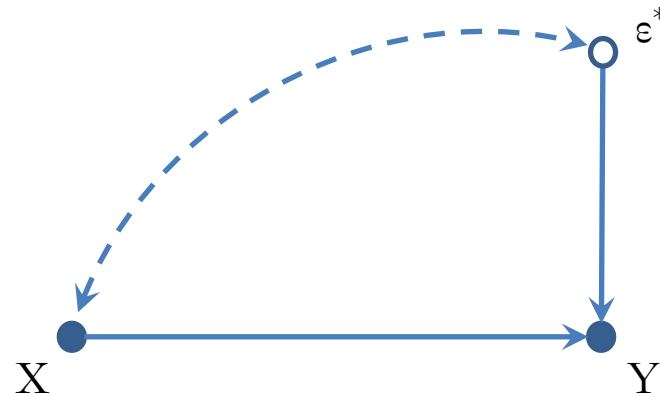
Multivariate
Extension, with
other Covariates



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ITCV information



“In settings where **valid instruments are not available**, the question arises how to evaluate OLS estimates...how large does the **endogeneity problem** have to be to make the coefficient statistically insignificant?”
(Larcker and Rusticus, 2010: 202)

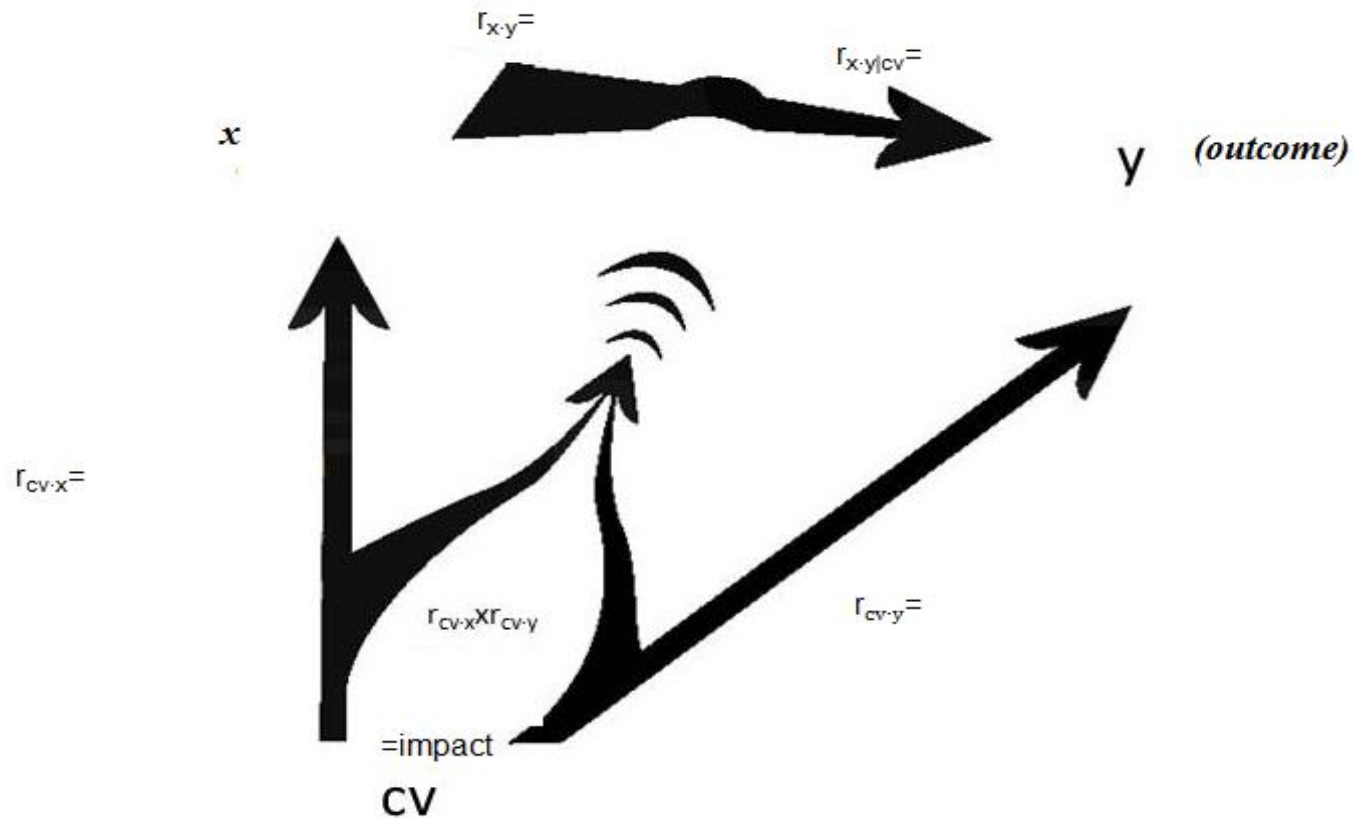


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ITCV Figure




Impact of a Confounding Variable on a Regression Coefficient



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Influence of Different Inputs on the ITCV

	Increased sample size	Larger Coefficient for IV	Larger SE for IV
Effect on the ITCV			
Interpretation of the effect on the ITCV	Increasing the sample size decreases the ITCV since effects are more power with larger samples. This means that it <i>more likely</i> a confounding variable exists that can invalidate statistical inference	A larger coefficient for the independent variable increases the ITCV since it move the effect size further away from zero. This means that it is <i>less likely</i> a confounding variable exists that can invalidate statistical inference	A larger standard error for the independent variable decreases the ITCV since it sampling distribution of the coefficient. This means that it <i>more likely</i> a confounding variable exists that can invalidate statistical inference



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ITCV Resources

- Frank (2000); Frank et al. (2013), Xu et al. (2019)
- Stata, SAS, R code for ITCV
- Kon-Found it! Excel sheet for ITCV
- Website:
 - <https://msu.edu/~kenfrank/research.htm>



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The relationship between IQ and Wages

RQ: What is the relationship between IQ and wages

- Key variables:
- DV – Log Wages
- IV – IQ Score
- Controls – Completed Year of Schooling (s), Experience in Years (expr), Tenure in years (tenure), Residency in Southern US (rns), Reside in Metro Area (smsa), and Year Fixed Effects



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Installing Stata Commands

```
ssc install konfound
```

```
ssc install moss
```

```
ssc install matsort
```

```
ssc install indeplist
```



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Summary Statistics

```
use http://www.stata-press.com/data/imeus/griliches, clear
```

```
sum lw s expr tenure rns smsa iq med kww age mrt, sep(0)
```

Variable	Obs	Mean	Std. dev.	Min	Max
-----+-----					
lw	758	5.686739	.4289494	4.605	7.051
s	758	13.40501	2.231828	9	18
expr	758	1.735429	2.105542	0	11.444
tenure	758	1.831135	1.67363	0	10
rns	758	.2691293	.4438001	0	1
smsa	758	.7044855	.456575	0	1
iq	758	103.8562	13.61867	54	145
med	758	10.91029	2.74112	0	18
kww	758	36.57388	7.302247	12	56
age	758	21.83509	2.981756	16	30
mrt	758	.5145119	.5001194	0	1



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Correlation

```
. corr lw s expr tenure rns smsa _I* iq
(obs=758)
```

	lw	s	expr	tenure	rns	smsa	_Iyea~67	_Iyea~68	_Iyea~69	_Iyea~70	_Iyea~71	_Iyea~73
lw	1.0000											
s	0.5027	1.0000										
expr	0.0846	-0.2418	1.0000									
tenure	0.1638	-0.0496	0.2307	1.0000								
rns	-0.1496	-0.0648	0.0058	-0.0366	1.0000							
smsa	0.2156	0.1021	-0.0332	0.0331	-0.1611	1.0000						
_Iyear_67	-0.1920	-0.1104	-0.0819	-0.0668	-0.0749	0.0065	1.0000					
_Iyear_68	-0.1261	-0.0794	-0.1842	-0.1230	-0.0220	-0.0062	-0.1027	1.0000				
_Iyear_69	0.0063	-0.0027	-0.1651	-0.1066	0.0012	-0.0447	-0.1070	-0.1212	1.0000			
_Iyear_70	0.0834	0.0704	-0.1541	0.0023	-0.0452	0.0511	-0.0914	-0.1036	-0.1079	1.0000		
_Iyear_71	0.1386	0.0955	0.1844	0.0496	-0.0069	-0.0426	-0.1119	-0.1268	-0.1321	-0.1129	1.0000	
_Iyear_73	0.3817	0.4398	0.0565	-0.0569	0.0548	0.0761	-0.1545	-0.1750	-0.1824	-0.1558	-0.1907	1.0000
iq	0.3471	0.5131	-0.1663	0.0194	-0.1339	0.0992	-0.0516	-0.0675	0.0087	0.1165	0.0698	0.2021



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Regression results

```
reg lw s expr tenure rns smsa _I* iq
```

Source	SS	df	MS	Number of obs	=	758
				F(12, 745)	=	46.86
Model	59.9127611	12	4.99273009	Prob > F	=	0.0000
Residual	79.3733888	745	.106541461	R-squared	=	0.4301
				Adj R-squared	=	0.4210
Total	139.28615	757	.183997556	Root MSE	=	.32641

lw	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
s	.0619548	.0072786	8.51	0.000	.0476658	.0762438
expr	.0308395	.0065101	4.74	0.000	.0180592	.0436198
tenure	.0421631	.0074812	5.64	0.000	.0274763	.0568498
rns	-.0962935	.0275467	-3.50	0.001	-.1503719	-.0422151
smsa	.1328993	.0265758	5.00	0.000	.0807268	.1850717
_Iyear_67	-.0542095	.0478522	-1.13	0.258	-.1481506	.0397317
_Iyear_68	.0805808	.0448951	1.79	0.073	-.0075551	.1687168
_Iyear_69	.2075915	.0438605	4.73	0.000	.1214867	.2936963
_Iyear_70	.2282237	.0487994	4.68	0.000	.132423	.3240245
_Iyear_71	.2226915	.0430952	5.17	0.000	.1380889	.307294
_Iyear_73	.3228747	.0406574	7.94	0.000	.2430579	.4026915
iq	.0027121	.0010314	2.63	0.009	.0006873	.0047369
_cons	4.235357	.1133489	37.37	0.000	4.012836	4.457878



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Stata konfound Output

```
. quietly: reg lw s expr tenure rns smsa _I* iq
```

```
. konfound iq
```

The Threshold for % Bias to Invalidate/Sustain the Inference

For iq:

To invalidate the inference 25.34% of the estimate would have to be due to bias; to invalidate the inference 25.34% (192) cases would have to be replaced with cases for which there is an effect of 0.

Impact Threshold for Omitted Variable

For iq:

An omitted variable would have to be correlated at 0.161 with the outcome and at 0.161 with the predictor of interest (conditioning on observed covariates) to invalidate an inference.
Correspondingly the impact of an omitted variable (as defined in Frank 2000) must be $0.161 \times 0.161 = 0.0260$ to invalidate an inference.

These thresholds can be compared with the impacts of observed covariates below.

Observed Impact Table for iq

Raw	Cor(v, X)	Cor(v, Y)	Impact
s	.5131	.5027	.258
_Iyear_73	.2021	.3817	.0771
smsa	.0992	.2156	.0214
rns	-.1339	-.1496	.02
_Iyear_67	-.0516	-.192	.0099
_Iyear_70	.1165	.0834	.0097
_Iyear_71	.0698	.1386	.0097
_Iyear_68	-.0675	-.1261	.0085
tenure	.0194	.1638	.0032
_Iyear_69	.0087	.0063	.0001
expr	-.1663	.0846	-.0141

Partial	Cor(v, X)	Cor(v, Y)	Impact
s	.4033	.3564	.1437
_Iyear_70	.0867	.1767	.0153
rns	-.1077	-.1373	.0148
tenure	.0603	.2071	.0125
_Iyear_71	.0544	.1904	.0104
_Iyear_73	.0302	.2809	.0085
smsa	.0339	.1826	.0062
_Iyear_69	.0306	.173	.0053
_Iyear_67	.0137	-.04	-.0005
_Iyear_68	-.0101	.0644	-.0006
expr	-.0608	.1649	-.01

X represents iq, Y represents lw, v represents each covariate.
First table is based on unconditional correlations, second table is based on partial correlations.



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Spreadsheet

User enters values in yellow

	A	B	C	D	E	F
1	KonFound-it! ©: BASICS					
2	unstandardized	User enters values in yellow				
3	estimated effect	standard error	sample size	# of covariates	α (significance level)	null hypothesis
4	0.0027121	0.0010314	758	12	0.05	0
5	For a 1-tailed test, double the size of α in cell E4					
6	Calculated Values			name of predictor of interest:		
7	df	default t critical (2-tailed)	t critical	IQ		
8	745	1.963	1.963	To override cell C8, type in your own value		
9	The default sign of t critical is the same as the sign of the estimated effect.					
10						
11	Publishable statements					
12	Replacement of Cases			Correlation Based (linear models only)		
13	To invalidate the inference 25% (192) of the cases would have			To invalidate the inference an omitted variable would have to be correlated		
14	to be replaced with cases for which there is an effect of zero.			at .161 with IQ and at .161 with the outcome, conditional on covariates.		
15						

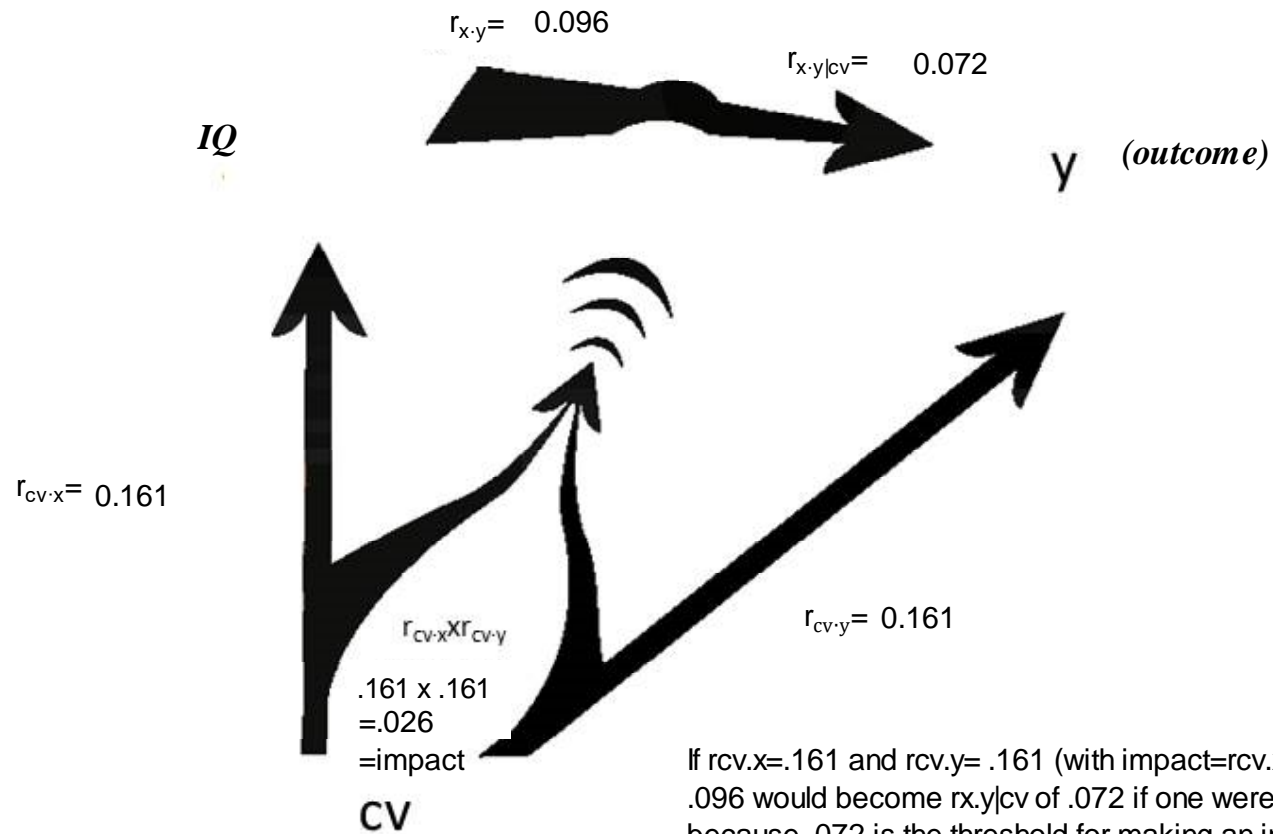


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Impact Figure

Impact of a Confounding Variable on a Regression Coefficient



(omitted confounding variable)

If $r_{cv,x} = .161$ and $r_{cv,y} = .161$ (with $\text{impact} = r_{cv,x} \times r_{cv,y} = .026$) then the $r_{x,y}$ of $.096$ would become $r_{x,y|cv}$ of $.072$ if one were to control for the confound (cv). Therefore because $.072$ is the threshold for making an inference, To invalidate the inference an omitted variable would have to be correlated at $.161$ with IQ and at $.161$ with the outcome, conditional on covariates.



[Click here to view for updates about upcoming workshops and other news through the KonFound-it! mailing list.](#)

KonFound-It! takes four values from many statistical analyses - the estimated effect (such as an unstandardized regression coefficient or the group mean difference), its standard error, the number of observations, and the number of covariates (and, for non-linear models, an additional value). KonFound-It returns output in the forms of publishable statements as well as figures to support the interpretation of the output.

Change or set any of the values below and then click run to see output from KonFound-It!

Estimated Effect

.0027121

Standard Error

.0010314

Number of Observations

758

Number of Covariates

12

Note that decimals must be denoted with a period, e.g., 2.1

RUN

Results (Printed)

Threshold Plot

Correlation Plot

Percent Bias Necessary to Invalidate the Inference:

To invalidate an inference, 25.342% of the estimate would have to be due to bias. This is based on a threshold of 0.002 for statistical significance ($\alpha = 0.05$). To invalidate an inference, 192 observations would have to be replaced with cases for which the effect is 0.

See Frank et al. (2013) for a description of the method

Citation: Frank, K.A., Maroulis, S., Duong, M., and Kelcey, B. 2013. What would it take to change an inference? Using Rubin's causal model to interpret the robustness of causal inferences. *Education, Evaluation and Policy Analysis*, 35 437-460.

Impact Threshold for a Confounding Variable:

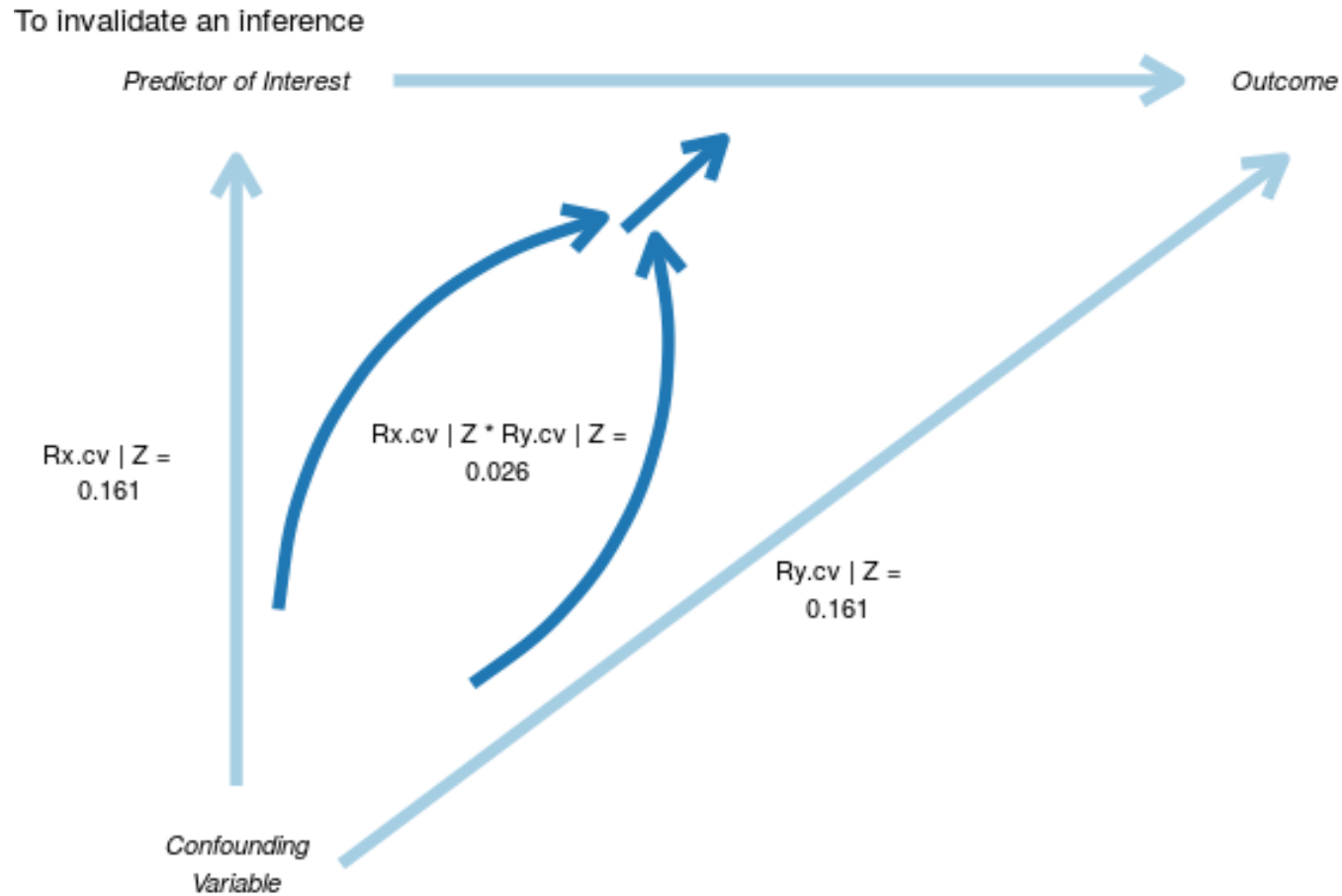
The minimum impact to invalidate an inference for a null hypothesis of 0 effect is based on a correlation of 0.161 with the outcome and at 0.161 with the predictor of interest (conditioning on observed covariates) based on a threshold of 0.072 for statistical significance ($\alpha = 0.05$). Correspondingly the impact of an omitted variable (as defined in Frank 2000) must be $0.161 \times 0.161 = 0.026$ to invalidate an inference for a null hypothesis of 0 effect.

See Frank (2000) for a description of the method

Citation: Frank, K. 2000. Impact of a confounding variable on the inference of a regression coefficient. *Sociological Methods and Research*, 29 (2), 147-194

- [Published empirical examples](#)
- [Full publishable write-up \(replacement of cases\)](#)
- [Full publishable write-up \(correlation\)](#)

Impact Figure from Konfound-it App



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Study on ITCV's application in management

- Is OVB as big of a problem in management research as extant scholarship suggests?
- Impact Threshold of a Confounding Variable in management journals
- Re-Examining the Semadeni et al. (2014) Simulation



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ITCV Values from a Content Analysis of Management Journals

		Panel A: Summary Results					Panel B: Covariates Higher than ITCV										Panel C: ITCV Value					
Journal Outlet	Number of Relationships	<u>Minimum correlation</u>		<u>Path correlation</u>			<u>Minimum correlation</u>					<u>Path correlation</u>										
		% Biased: One Covariate	% Biased: Two Covariates	% Biased: One Covariate	% Biased: Two Covariates		10%	25%	Median	75%	90%	10%	25%	Median	75%	90%						
AMJ	130	26.15%	13.08%	39.23%	26.15%		0	0	0	1	2	0	0	0	2	3	0.040	0.134	0.238	0.415	0.609	0.286
JAP	94	20.21%	11.70%	29.79%	19.15%		0	0	0	0	2	0	0	0	1	2	0.126	0.191	0.345	0.483	0.671	0.367
JOM	47	21.28%	12.77%	38.30%	21.28%		0	0	0	0	2	0	0	0	1	3	0.058	0.193	0.274	0.407	0.603	0.323
SMJ	111	35.14%	23.42%	48.65%	34.23%		0	0	0	1	3	0	0	0	2	4	0.042	0.085	0.148	0.295	0.463	0.211
Overall	382	26.70%	15.71%	39.53%	26.18%		0	0	0	1	2	0	0	0	2	3	0.046	0.132	0.245	0.400	0.597	0.288
Macro	172	33.14%	20.35%	46.51%	31.98%		0	0	0	1	3	0	0	0	2	4	0.045	0.096	0.171	0.308	0.494	0.232
Micro	210	21.43%	11.90%	33.81%	21.43%		0	0	0	0	2	0	0	0	1	3	0.067	0.176	0.310	0.460	0.638	0.335

Notes: “% Biased: One Covariate” and “% Biased: Two Covariates” reflect the proportion of total relationships that are potentially biased if at least one or two control variables, respectively, features the properties of correlations higher than the ITCV value.



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Revisiting the Simulation

- Original simulation – Semadeni et al. (2014)
 - Correlated the independent variable with the error term
 - 0.10 = “Low endogeneity”
 - 0.30 = “Moderate endogeneity”
 - Error term represents about 90% of the variance
 - Vector of omitted variables or REALLY strong omitted variable

$$y = \alpha + \beta_{1 \times 1} + e$$

$$\text{Corr}[x, e] = 0.1 \text{ or } 0.3$$

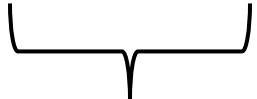


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Revisiting the Simulation

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 - Vector of omitted variables or REALLY strong omitted variable

$$y = \alpha + \beta_{1 \times 1} + \beta_{2 \times 2} + e$$


Corr[x1,x2] =

- 0.05 (10th percentile)
- 0.10 (25th percentile)
- 0.15 (50th percentile)
- 0.25 (75th percentile)
- 0.40 (90th percentile)



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Simulation Results

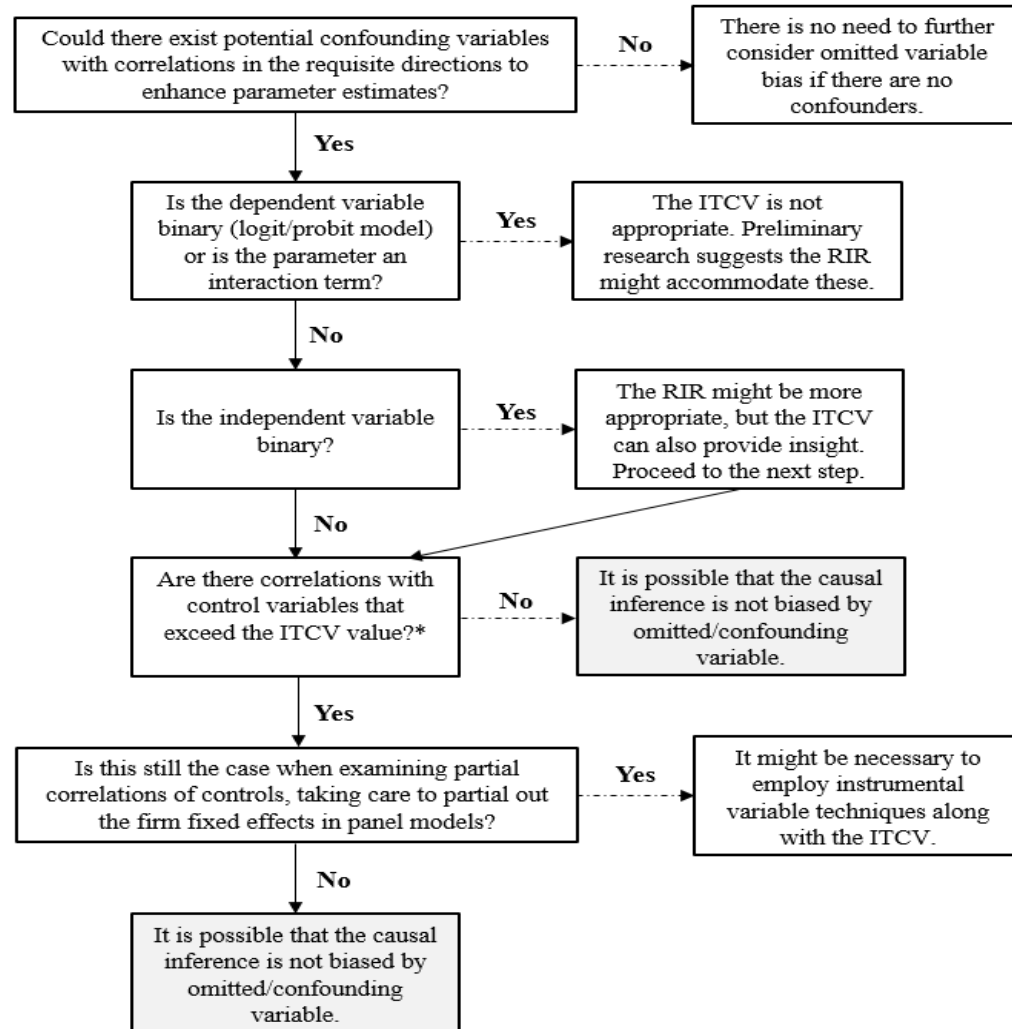
Regression Model	Median	10th Pctile	25th Pctile	75th Pctile	90th Pctile	StDev.
Panel A (r=0.00)						
OLS (Fully Specified)	0.127	0.057	0.089	0.162	0.191	0.053
OLS (Not Fully Specified)	0.126	0.058	0.089	0.163	0.189	0.053
2SLS w/ Weak Instruments	0.101	-1.882	-0.710	0.871	1.884	4.104
2SLS w/ Moderate Instruments	0.133	-0.178	-0.023	0.279	0.436	0.239
2SLS w/ Strong Instruments	0.128	0.009	0.071	0.185	0.239	0.091
Panel C (r=0.10)						
OLS (Fully Specified)	0.122	0.057	0.088	0.159	0.193	0.053
OLS (Not Fully Specified)	0.137	0.072	0.103	0.172	0.206	0.053
2SLS w/ Weak Instruments	0.071	-1.581	-0.604	0.906	2.102	2.886
2SLS w/ Moderate Instruments	0.129	-0.170	-0.032	0.295	0.451	0.251
2SLS w/ Strong Instruments	0.120	0.003	0.054	0.180	0.232	0.089
Panel E (r=0.25)						
OLS (Fully Specified)	0.129	0.056	0.092	0.165	0.196	0.054
OLS (Not Fully Specified)	0.163	0.092	0.126	0.195	0.227	0.052
2SLS w/ Weak Instruments	0.113	-1.732	-0.719	0.855	2.075	2.580
2SLS w/ Moderate Instruments	0.108	-0.196	-0.045	0.274	0.422	0.245
2SLS w/ Strong Instruments	0.125	0.020	0.073	0.189	0.245	0.088
Panel F (r=0.40)						
OLS (Fully Specified)	0.127	0.051	0.088	0.165	0.200	0.059
OLS (Not Fully Specified)	0.181	0.111	0.146	0.217	0.245	0.052
2SLS w/ Weak Instruments	0.239	-1.750	-0.586	1.009	2.035	3.318
2SLS w/ Moderate Instruments	0.129	-0.168	-0.022	0.291	0.429	0.235
2SLS w/ Strong Instruments	0.127	0.011	0.065	0.188	0.242	0.091



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Decision Tree of When to Employ the ITCV



- It is important to reiterate that the ITCV represents the square root of the product of correlations between a potential omitted variable and both the independent and dependent variables. This is therefore the case when examining control variables as potential proxies for an omitted variable. Specifically, it is essential to compare the square root $\text{corr}[\text{control}, y] \times \text{corr}[\text{control}, x]$ against the ultimate ITCV value.



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Limitations of the ITCV

- The focus is on causal inference, not effect sizes
 - Provides information about statistical significance, not coefficient interpretation
 - Endogeneity can still bias the coefficients
 - Relies on an understanding of the actual DGP (true population model)
- IV techniques may necessary if translated effect sizes matter



Limitations of the ITCV (cont.)

- A confounder may exist that exhibits a correlation with the IV and DV at values greater than control covariates
- The appropriateness of the ITCV also depends on the nature of the empirical estimation procedure and corresponding data
- The ITCV is also currently unable to address interaction terms because the marginal effects of the relationships are contingent on the values of the lower order constituents



Summary

- Confounders and OVB are a major concern in organizational research
- Instrumental Variables Techniques can address the issue, but key assumptions must be met
- The ITCV may help to understand how statistical inferences change because of OVB



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